ESTIMATING TREE BOLE HEIGHT WITH BAYESIAN ANALYSIS

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ABSTRACT. The classical method for estimating a height-diameter model is based on the Least Squares Method (LSM) and the fit of a regression line. The Bayesian method has an exclusive advantage, compared with the classical method, in that the parameters to be estimated are considered as random variables. In this study, the Simple Linear Regression (SLR) model and the Bayesian model were used to estimate bole height from breast height diameter. We used data of the forest stands of Rhodope (north-eastern Greece). variables that we used were the tree bole height and the diameter at breast height. The results showed that there is an improvement in prediction accuracy with the Bayesian model; however, this didn't lead to narrower confidence intervals of the predicted value, compared to SLR. Narrower confidence intervals are not necessarily achieved with Bayesian methods; confidence intervals' width is related to both statistical analyses and nature of data (in this case, species ecology and structure - composition of the stands where the sampled trees belong).

Keywords: Bayes, regression; tree height; tree bole.

1 INTRODUCTION

Forests play an important role in the wood production, recreation, carbon sinks and climate change (FAO 2006). Height and diameter are the most common variables measured to estimate tree volume, site quality and other important variables in forest growth and yield, species' succession and carbon estimation models (Peng et al. 2001). Breast height diameter of individual trees is easy to measure with high accuracy in the field, and minimum cost. On the contrary, tree height is not easy to estimate on standing trees; it is a timeconsuming procedure, subjected to observation error, and affected by optical obstacles (Colber et al. 2002). Tree volume estimation and site quality, as well as the description of a stand's dynamics and species' succession over time, are greatly dependent by the accuracy of height-diameter models (Curtis 1967). There is a series of height-diameter models, developed for several forest species (Peng et al. 2004, Temesgen et al. 2007, Van Laar & Akca 2007). These models can be used to estimate a missing height for a tree, while having the diameter at breast height measured (Shongming et al. 1992, Hann 2006), to estimate indirectly the height increment (Larsen & Hann 1987), and to estimate tree biomass applying biomass models (Penner et al. 1997). Chave et al. (2005) found that the most important parameters in biomass estimation of tropical forests, in decreased order, are the breast height diameter, the wood density, the height and the forest's moisture (dry forest, forest with medium or heavy moisture). Comprising height as a variable, in biomass models, reduces significantly the standard error of biomass estimation. Therefore, an accurate tree height estimation should be imposed in forest inventories, simulation models, forest management and decision support systems (Peng et al. 2001, Colber et al. 2002, Curtis 1967).

Curtis (1967) compared the fit of linear heightdiameter models, in data of Douglas-fir trees (*Pseudot*suga menziesii (Mirb.) Franco). Since then, with the relatively easy fit of nonlinear models, many nonlinear models have been developed for estimating height (Fekedulegn et al. 1999). However, because the tree shape and allometry are affected by competition and environmental factors (Batziou et al. 2016, King 1991, Kitikidou et al. 2016), any changes in these conditions over time will probably affect the height-diameter relationship. This can cause uncertainty in height estimation from the diameter. An important limitation in these models is that they produce very different results, when applied in different stands from those where they originally were developed (Calama & Montero 2004, Nogueira et al. 2008). Also, the height-diameter relationship is not stable over time, even in the same stand (Flewelling & Jong 1994, Lappi 1991). These differences could have significant effects in biomass and dynamic carbon sinks. Uncertainty in height estimation, due to time changes, must be taken into account, when heightdiameter relationships in natural stands are examined, and there are no available methods to resolve this problem.

From all these points previously described, one can assess that, even a small improvement in height-diameter models, is worthy of investigation, especially regarding bole height, which is referring to merchantable vol-Merchantable volume is the main interest of ume. forest inventories and the most contributing variable in forest biomass and carbon sinks. Bayesian analysis is an alternative method of statistical inference, frequently used in ecological models' evaluation (Anholt et al. 2000, Toivonen et al. 2001, Shen et al. 2003). In forestry, Bayesian methods were adopted in several applications, like the estimation of aboveground biomass (Zapata-Cuartas et al. 2012), in diameter distribution (Bullock & Boone 2007) and basal area distribution (Nyström & Ståhl 2001), in tree growth estimation (Clark et al. 2007), mortality estimation (Wyckoff & Clark 2000, Metcalf et al. 2009), stands' height and volume estimation (Stewart & Weiskittel 2012), and in individual tree height estimation (Zhang et al. 2014).

Fagus sylvatica L. expands into a wide area of Europe and up to southern Scandinavia, but in the Mediterranean area appears only in the mountains (Korakis 2015). Fagus sylvatica is a very important tree species for Greece, since beech forests compose 5.17% of Greek forests, while beech stands provide 20.05% of the merchantable volume of forests in the country (Ministry of Agriculture 1992).

In this study, we have developed Bayesian heightdiameter models for beech trees, examining numerous trials from one wide study area. Since nonlinearity in a height-diameter relationship is not always detectable within an even-aged stand, due to sample sizes that cannot display lack of fit of the linear model, or excessive random variability of heights within certain diameter classes (Van Laar & Akca 2007), Simple Linear Regression (SLR) models were developed. We compared the Bayesian height-diameter models to their corresponding SLR models (i.e. regression models developed with Least Squares Method (LSM)), in order to see if there is any improvement in height estimation.

2 MATERIALS AND METHODS

2.1 Study area

The study was conducted in the central part of the Rhodope mountains, in the north part of the Xanthi region, in north-eastern Greece (Fig. 1). The climate can be characterized as humid with harsh winters. The soils in the areas where data was collected are mainly acid brown forest soils (Dystric Cambisoils) (Milios 2000). The data were taken from areas having an elevation that ranges from approximately 580 to 1700 m.



Figure 1: Study area (wide view).

The elevation of the western part of the study area (Fig. 2, red polygon points) ranges from 1100 to 1725 m, while that of the eastern part (Fig. 2, green polygon points) ranges from 580 to 750 m.



Figure 2: Study area (large display).

In the total study area, 2809 *F. sylvatica* trees were measured in plots of 500 m² (25 m x 20 m). The plots had been established randomly, in the context of other

Trial No.	Trees	Variable	Mean	Standard	Min.	Max.
(Productivity)	No. (n)			Deviation		
1	2809	D (cm)	14.78	10.95	4	67
1		Hb (m)	3.65	3.63	0	23.5
9	1405	D(cm)	12.4	8.75	4	67
2	1495	Hb (m)	3.02	2.43	0	16
9	1900	D(cm)	13.4	10.32	4	63
0	1598	Hb(m)	3.26	3.13	0	20
4	2157	D(cm)	13.48	9.99	4	67
4		Hb (m)	3.3	2.95	0	20
F	759	D(cm)	13.48	9.28	4	62
9	199	Hb (m)	2.39	2.42	0	14.5
C	910	D (cm)	27.85	13.72	4	65
0	210	Hb (m)	8.98	6.31	1	23.5
7	1596	D(cm)	14.24	11	4	63
1	1520	Hb (m)	3.53	3.45	0	20
8	2367	D (cm)	14.75	11.15	4	67
		Hb (m)	3.8	3.75	0	23.5
0	754	D(cm)	16.69	11.56	4	67
9		Hb (m)	4.7	4.24	0.5	23.5
10	1968	D (cm)	14.39	10.75	4	63
10		Hb (m)	3.38	3.33	0	20
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Table 1: Descriptive statistics of measured variables

studies, for the analysis of stand structure and dynamics in pure and mixed stands in areas of various productivities. In each tree, the breast height diameter and the bole height were measured. In particular, the measured trees came from pure beech stands in good and medium productivity sites, from Pinus sylvestris - F. sylvatica stands growing in good, medium and poor productivity sites, from Fagus sylvatica – Abies borisii-regis stands growing in medium productivity sites, and from Quercus petraea – Fagus sylvatica found in medium productivity sites (Milios 2000). In P. sylvestris – F. sylvatica stands, P. sylvestris was the dominant species of the overstorey in almost all cases (Milios 2000). From different mixtures and productivity sites, ten different combinations (trials) were created for the analysis. Descriptive statistics of measured variables, in each trial, are given in Table 1.

2.2 Trial No (Productivity)

1. All

2. Good for Pinus sylvestris - Fagus sylvatica, medium for Pinus sylvestris - Fagus sylvatica, and poor for Pinus sylvestris - Fagus sylvatica

- 3. Medium for *Pinus sylvestris Fagus sylvatica*, medium for *Fagus sylvatica - Abies borisii-regis*, and medium for *Quercus petraea - Fagus sylvatica* (western part)
- 4. Good for Pinus sylvestris Fagus sylvatica, medium for Pinus sylvestris - Fagus sylvatica, poor for Pinus sylvestris - Fagus sylvatica, medium for Fagus sylvatica - Abies borisii-regis, and medium for Quercus petraea - Fagus sylvatica (western part)
- 5. Medium for *Quercus petraea* Fagus sylvatica (western part), and medium for *Quercus petraea* -Fagus sylvatica (eastern part)
- 6. Good for Fagus sylvatica, and medium for Fagus sylvatica
- 7. Medium for *Pinus sylvestris Fagus sylvatica*, medium for *Fagus sylvatica* - *Abies borisii-regis*, medium for *Quercus petraea* - *Fagus sylvatica* (western part), and medium for *Fagys sylvatica*
- 8. Good for Pinus sylvestris Fagus sylvatica, medium for Pinus sylvestris - Fagus sylvatica, medium for Fagus sylvatica - Abies borisii-regis, medium for Quercus petraea - Fagus sylvatica (western part), good for Fagus sylvatica, and medium for Fagus sylvatica

9. Good

10. Medium

2.3 SLR models developed with LSM

In a scatterplot with an independent (X) variable and a dependent (Y) variable, the goal of a linear regression model is to fit a line through the points. Specifically, with LSM, the squared deviations of the observed points from that line are minimized.

2.4 Bayesian model

Suppose $y = (y_1, y_2, y_3, ...)$ is a vector of data and $\theta = (\theta_1, \theta_2, \theta_3, ...)$ is a vector of parameters, which will be estimated. The Bayes rule is expressed as follows:

$$p(y,\theta) = p(y|\theta) p(\theta) = p(\theta|y) p(y)$$
(1)

where p is a density probability function. While the values of θ are estimated with LSM or Maximum Likelihood Estimation in the classical approach, in the Bayesian approach we use probability distributions to describe the uncertainty in the parameters that are about to be estimated. The θ have a probability distribution, which is calculated as another form of (1):

$$p(\theta | y) = \frac{p(y | \theta) p(\theta)}{p(y)}$$
(2)

where $p(y) = \int p(y|\theta) p(\theta) d\theta$ for continuous θ . The integration of the acceptable values θ , p(y) is not dependent on θ and can be considered as constant for constant y, a fact that leads to the relationship (3):

$$p(\theta | y) \propto p(y | \theta) p(\theta)$$
(3)

What we are interested in the Bayesian analysis is the estimation of the conditional probability, i.e., the posterior probability distribution. The $p(y|\theta)$ is giving us the distribution of y assuming that the θ is known, i.e., it is the function of maximum likelihood when it is considered as a function of the θ parameters (Edwards 1992). The $p(\theta)$ is the prior probability distribution of the θ parameters, and it represents all available information regarding y. So, the relation (3) is suggesting that the posterior distribution of θ is analogous to the likelihood of y, given the θ and the prior distribution of θ .

The important characteristic of the Bayesian analysis is that the models' parameters are considered to be random variables (Stewart & Weiskittel 2012), while in the classical method of Least Squares, the models' parameters are considered fixed values (Edwards 1992, De Valpine & Hastings 2002).

The selection of the prior distribution is essential in the Bayesian analysis (Gelman et al. 2004). If there is no available information regarding parameters' distribution, we can accept ignorance of the prior distribution, i.e. accept an uninformative, Gaussian prior distribution.

Bayesian parameters are estimated by applying the SPSS AMOS (Analysis of Moment Structures) software, v.21.0 (Arbuckle 2012). We defined to run 100000 iterations, from which the initial 500 were considered as initial stages of the chain before it converges (burn-in period).

2.5 Comparison of the classical and the Bayesian method

Three statistical criteria were used for the classical (SLR - LSM) and the Bayesian method comparison (Ki-tikidou 2005):

- Absolute mean error $Bias = \frac{\sum_{i=1}^{n} |H_{b_i} \hat{H}_{b_i}|}{n}$ (optimum value = 0).
- Standard error of the estimate of theoretical values $se = \sqrt{\frac{\sum_{i=1}^{n} (H_{b_i} \hat{H}_{b_i})^2}{n-p}} \text{ (optimum value = min).}$
- Coefficient of determination $R^2 = 1 \frac{\sum\limits_{i=1}^{n} (H_{b_i} \hat{H}_{b_i})^2}{\sum\limits_{i=1}^{n} (H_{b_i} \overline{H_b})^2}$ (optimum value = 1).

where:

 H_{bi} = bole height of the *i*-th tree \hat{H}_{bi} = estimated bole height of the *i*-th tree \overline{H}_{b} = mean bole height of the sampled trees n = sample size (number of trees in each trial) p = number of the model's parameters = 2.

3 Results

From the ten trials of Table 1, bole height estimation was improved by applying the Bayesian method, in all of them. Comparison statistics for these ten trials, the estimated mean bole height, its confidence interval, and variance, are shown in Table 2, while SLR are shown in Table 3. We observe that in general (i.e. except for trials 2, 6, and 9), the Bayesian models gave bigger values for the estimated mean bole height, wider confidence intervals, and bigger variance.

In the trace plots of Figures 3a-3j, the convergence that was succeeded by running the algorithm, is assessed. The fast, up-and-down interchange, without trends (patterns), shows that the convergence was succeeded in a short time.



Figure 3: Trace plot of the variance of the breast height diameter D (a: Trial 1, b: Trial 2, c: Trial 3, d: Trial 4, e: Trial 5, f: Trial 6, g: Trial 7, h: Trial 8, i: Trial 9, j: Trial 10).

Table 2: Statistics for the trials in which bole height estimation was improved with the Bayesian method (SE – Standard Error; LB – Lower Bound; UB – Upper Bound; CI – Confidence Interval; EM – Estimated Mean; No – Number; ht – height).

Trial	No of		Classical method (SLR, LSM)					Bayesian method					
No	trees	Bias	SE	R2	$\mathbf{E}\mathbf{M}$	CI of	the EM	Bias	SE	R2	$\mathbf{E}\mathbf{M}$	CI of	the EM
	(n)				$\mathrm{bole}\mathrm{ht}$	LB	UB				$\mathrm{bole}\mathrm{ht}$	LB	UB
1	2809	2.103	3.009	0.314	3.6538	1.654	13.348	1.706	2.5702	0.499	3.8124	0.642	24.727
2	1495	1.593	2.1891	0.191	3.0204	1.999	9.6609	1.382	2.1184	0.243	2.9807	0.584	21.284
3	1398	1.756	2.5304	0.347	3.2614	1.582	12.12	1.557	2.4088	0.408	3.5916	0.73	24.595
4	2157	1.798	2.51	0.276	3.2976	1.827	11.599	1.566	2.393	0.342	3.4607	0.642	24.605
5	753	1.558	2.2673	0.126	2.3923	1.512	6.901	1.331	2.0033	0.318	2.8157	0.554	17.051
6	210	4.286	5.2663	0.306	8.9767	2.91	18.426	3.77	4.7892	0.426	5.2725	0.233	19.358
7	1526	1.898	2.7464	0.367	3.5265	1.579	12.796	1.644	2.5164	0.469	3.8676	0.731	24.728
8	2367	2.096	3.0081	0.356	3.8015	1.644	14.285	1.726	2.6168	0.513	3.9402	0.665	24.728
9	754	2.589	3.5377	0.304	4.7031	2.139	14.872	1.954	2.8066	0.562	4.1765	0.598	22.972
10	1968	1.953	2.7852	0.3	3.3775	1.619	11.607	1.633	2.4648	0.451	3.7044	0.652	24.733

In Figures 4a-4j, the autocorrelation of the values of the variance of D, during iterations, is illustrated. The lag across the horizontal axis is referred to the interval in which the autocorrelation is estimated. In common situations, we expect that the autocorrelation coefficient

is reducing till zero, and stays close to zero in all the following lag intervals. As illustrated on Figures 4a-4j, the initial 500 iterations (burn-in period) were more than enough, to ensure convergence.



Figure 4: Autocorrelation of the variance of the breast height diameter D (a: Trial 1, b: Trial 2, c: Trial 3, d: Trial 4, e: Trial 5, f: Trial 6, g: Trial 7, h: Trial 8, i: Trial 9, j: Trial 10).

Table 3: SLR models for each trial.

Trial No	Classical method (SLR, LSM)
1	$\hat{H}_b = 0.911 + 0.186D$
2	$\hat{H}_b = 1.512 + 0.122D$
3	$\hat{H}_b = 0.868 + 0.179D$
4	$\hat{H}_b = 1.207 + 0.155D$
5	$\hat{H}_b = 1.140 + 0.093D$
6	$\hat{H}_b = 1.892 + 0.254D$
7	$\hat{H}_b = 0.819 + 0.190D$
8	$\hat{H}_b = 0.841 + 0.201D$
9	$\hat{H}_b = 1.331 + 0.202D$
10	$\hat{H}_b = 0.942 + 0.169D$

In Figures 5a-5j, the classical (SLR, LSM) vs. the Bayesian model is illustrated, for each trial, in which the Bayesian model improved the bole height estimation.

4 DISCUSSION AND CONCLUSIONS

Estimating bole height from breast height diameter, applying the Bayesian method to the classical regression method (SLR, LSM) lead to an improvement of estimation accuracy (an increase of the coefficient of determination \mathbb{R}^2), in all our trials, conducted in the same wide study area. However, the confidence intervals of the estimated mean bole height were larger, a fact that is in contradiction to the study of Zhang et al. (2014). This is the result of the combination of the species ecology and the structure - composition of the stands where the measured trees came from. Beech is a shade tolerant species having growth plasticity (Assman 1970). As a result, in mixed stands where beech grows together with less shade tolerant species that dominate in the overstorey, beech trees in the understory and middle story have enough light to grow a long live crown. Shade intolerant species do not have as dense crowns as shade tolerant species, and they generate less shade, since, in contrast



Figure 5: Height-diameter model (a: Trial 1, b: Trial 2, c: Trial 3, d: Trial 4, e: Trial 5, f: Trial 6, g: Trial 7, h: Trial 8, i: Trial 9, j: Trial 10).

to shade tolerant species, after their youth, they cannot form shade leaves (see Dafis 1986, Oliver & Larson 1996, Lacointe et al. 2004). Thus, when shade-tolerant trees with large diameters of the overstorey, grow with shade intolerant species, they grow longer crowns compared to those growing in pure stands, or mixed stands having other shade tolerant species since more light reaches the lower part of boles. In the present study, this is obvious in all the trials where beech grows with *P. sylvestris* or Q. petraea. Pinus sylvestris is shade intolerant, while Q. petraea is moderately shade-intolerant (Korakis 2015). In these trials, the Bayesian method overestimates the bole height of trees having large diameters, compared to the classical regression method, leading to larger confidence intervals (Figs. 5a-5e, Figs. 5g-5j). Only in trial 6, which is the only trial with pure beech stands (and crowns of the trees with large diameters are short) the Bayesian method does not overestimate the bole height (Fig. 5f). So, the Bayesian method does not give a priori narrower confidence intervals; confidence intervals' width is related to the characteristics of the data.

Another finding of the present study is that in both methods (classical and Bayesian), in trials from mixed stands of P. sylvestris - F. sylvatica and/or Q. petraea - F. sylvatica (trials 2 and 5) the \mathbb{R}^2 values are lower than in the other trials. One should note that, in trials with pure beech stands (1, 6, 7, 8, 9 and 10) the \mathbb{R}^2 values in the Bayesian method are rather high, while in some of these trials (1, 8, 9 and 10) the increase of the \mathbb{R}^2 , compared to the classical method, is very high. The lowest \mathbb{R}^2 value of these trials is found in trial 6, where trees only from pure stands participate. This pattern indicates that bole height prediction is more accurate where trees from all growth environments are present. However, even in the case where only trees from pure beech stands are present, the Bayesian method provides a rather high \mathbb{R}^2 value. The behaviour of these models is affected by the growth characteristics and ecology of beech, which have been described previously. Probably the existence of trees from pure beech stands provides an adequate number of large trees having high bole height in the sample (small crown length is a result of competition among shade-tolerant trees), improving the model's performance.

In trial 9 (good productivity sites) the R^2 value of 0.56 is an indication that growth conditions in different site productivities may influence the model's performance. However, in trial 10 (medium productivity sites) the R^2 value was lower than that of trial 1.

To sum up, the Bayesian method is an important tool, used more and more by ecologists (Hui et al. 2006, Hui et al. 2011, McCarthy et al. 2007). For this specific application in bole height estimation, more independent variables could be added in the future, such as site quality index, age, or stand density (Temesgen & Gadow 2004, Newton & Amponsah 2007).

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References

- Anholt, B., Werner, E., & Skelly, D. 2000. Effect of food and predators on the activity of four larval ranid frogs, Ecology, 81 (12): 3509–3521.
- Arbuckle, J. 2012. IBM SPSS Amos 21 User's Guide, IBM Corporation, 653 p., USA.
- Assmann, E. 1970. The Principles of Forest Yield Study, Pergamon Press, 506 p., New York, USA.
- Batziou, M., Milios, E., & Kitikidou, K. 2016. Is diameter at the base of root collar a key characteristic of seedling sprouts in a *Quercus pubescens - Quercus frainetto* grazed forest in Northeastern Greece? A morphological analysis, New Forests, 48(1): 1–16.
- Bullock, B., & Boone, E. 2007. Deriving tree diameter distributions using Bayesian model averaging, Forest Ecology and Management, 242, (2–3): 127–132.
- Calama, R., & Montero, G. 2004. Interregional nonlinear height-diameter model with random coefficients for stone pine in Spain, Canadian Journal of Forest Research, 34(1): 150–163.
- Chave, J., Andalo, C., Brown, S., Cairns. M., Chambers, J., Eamus, D., Fölster, H., Fromard, F., Higuchi, N., Kira, T., Lescure, J., Nelson, B., Ogawa, H., Puig, H., Rira, B., & Yamakura T. 2005. Tree allometry and improved estimation of carbon stocks and balance in tropical forests, Oecologia, 145(1): 87–99.
- Clark, J. S., Wolosin, M., Dietze, M., Ibáñez, I., LaDeau, S., Welsh, M., & Kloeppel, B. 2007. Tree growth inference and prediction from diameter censuses and ring widths. Ecological Applications 17:19421953.
- Colbert, K., Larsen, D., & Lootens, J. 2002. Heightdiameter equations for thirteen midwestern bottomland hardwood species., Northern Journal of Applied Forestry, 19(4): 171–176.
- Curtis, R. 1967. Height-diameter and height-diameterage equations for second-growth Douglas-fir, Forest Science, 13: 365–375.
- Dafis, S. 1986. Forest Ecology (in Greek), Giahoudi-Giapouli, 443 p., Thessaloniki, Greece.

- De Valpine, P., & Hastings, A. 2002. Fitting population models incorporating process noise and observation error, Ecological Monographs, 72(1): 57–76.
- Edwards, A. 1992. Likelihood. Johns Hopkins University Press, 296 p, Baltimore, USA.
- FAO. 2006. GreenFacts Scientific facts on forests. Level2 Details on forests, GreenFacts Foundation, 62 p,Brussels, Belgium.
- Fekedulgn, D., Mac Siurtain, M., & Colbert, J. 1999. Parameter estimation of nonlinear growth models in forestry, Silva Fennica, 33(4): 327–336.
- Flewelling, J., & De Jong, R. 1994. Considerations in simultaneous curve fitting for repeated height-diameter measurements, Canadian Journal of Forest Research, 24(7): 1408–1414.
- Gelman, A., Carlin, J., Stern, H., & Rubin, D. 2004. Bayesian Data Analysis, 2nd edition. Chapman and Hall/CRC, 690 p, Boca Raton, USA.
- Hann, D. 2006. ORGANON User's Manual Edition 8.0. Department of Forest Resources, Oregon State University, 129 p, Corvallis, Oregon, USA.
- Huang, S. 1999. Ecoregion-based individual tree heightdiameter models for lodgepole pine in Alberta, Western Journal of Applied Forestry, 14(4): 186–193.
- Huang, S., Titus, S., & Wiens, D. 1992:. Comparison of nonlinear height-diameter functions for major Alberta tree species, Canadian Journal of Forest Research, 22(9): 1297–1304.
- Hui, C., Foxcroft, L., Richardson, D., & Macfadyen, S. 2011. Defining optimal sampling effort for largescale monitoring of invasive alien plants: a Bayesian method for estimating abundance and distribution, Journal of Applied Ecology, 48(3): 768–776.
- Hui, C., McGeoch, M., & Warren, M. 2006. A spatially explicit approach to estimating species occupancy and spatial correlation, Journal of Animal Ecology, 75(1): 140–147.
- King, D. 1991. Tree allometry, leaf size and adult tree size in old-growth forests, Tree Physiology, 9(3): 369– 381.
- Kitikidou, K. 2005. Applied Statistics Using SPSS, Tziolas Press, 288 p, Thessaloniki, Greece.
- Kitikidou, K., Milios, E., & Katsogridakis, S. 2016. Meta-analysis for the volume of *Pinus sylvestris* in Europe, Revista Chapingo Serie Ciencias Forestales y del Ambiente, 23(1): 22–34.

- Korakis, G. 2015. Forest Botany, Hellenic Academic Libraries, 619 p, Athens, Greece.
- Lacointe, A., Deleens, E., Ameglio, T., Saint-Joanis, B., Lelarge, C., Vandame, M., Song, G., & Daudet, F. 2004. Testing the branch autonomy theory: a 13C/14C double-labelling experiment on differentially shaded branches, Plant, Cell and Environment, 27: 1159–1168.
- Lappi, J. 1991. Calibration of height and volume equations with random parameters, Forest Science, 37: 781–801.
- Larsen, D., & Hann, D. 1987. Height-Diameter Equations for Seventeen Tree Species in Southwest Oregon. Oregon State University, 16 p, Corvallis, Oregon, USA.
- Li, R., Stewart, B., & Weiskittel, A. 2012. A Bayesian approach for modelling non-linear longitudinal/hierarchical data with random effects in forestry, Forestry, 85(1): 17–25.
- McCarthy, M. 2007. Bayesian Methods for Ecology, Cambridge University Press, 295 p, Cambridge, UK.
- Metcalf, C., McMahon, S., & Clark, J. 2009. Overcoming data sparseness and parametric constraints in modeling of tree mortality: a new nonparametric Bayesian model, Canadian Journal of Forest Research, 39(9): 1677–1687.
- Milios, E. 2000. Dynamics and evaluation of the Rhodope mixed stands in the region of Xanthi (In Greek), PhD thesis, Aristotle University of Thessaloniki, 345 p.
- Milios, E. 2004. The influence of stand development process on the height and volume growth of dominant Fagus sylvatica L. sl trees in the central Rhodope Mountains of north-eastern Greece, Forestry, 77(1): 17–26.
- Ministry of Agriculture. 1992. Results of the rst national forest inventory (in Greek), Publication of Department of Forest Mapping, 134 p, Athens, Greece.
- Newton, P., & Amponsah, I. 2007. Comparative evaluation of five height-diameter models developed for black spruce and jack pine stand-types in terms of goodnessof-fit, lack-of-fit and predictive ability, Forest Ecology and Management, 247(1–3): 149–166.
- Nogueira, E., Nelson, B., Fearnside, P., Frana, M., & Oliveira, Á. 2008. Tree height in Brazil's "arc of deforestation": shorter trees in south and southwest Amazonia imply lower biomass, Forest Ecology and Management, 255(7): 2963–2972.

- Nyström, K., & Ståhl, G. 2001. Forecasting probability distributions of forest yield allowing for a Bayesian approach to management planning, Silva Fennica, 35(2): 185–201.
- Oliver, C., & Larson, B. 1996. Forest stand dynamics, 520 p, John Wiley & Sons, Inc., New York, USA.
- Peng, C., Zhang, L., & Liu, J. 2001. Developing and validating nonlinear height-diameter models for major tree species of Ontario's boreal forests, Northern Journal of Applied Forestry, 18(3): 87–94.
- Peng, C., Zhang, L., Zhou, X., Dang, Q., & Huang, S. 2004. Developing and evaluating tree height-diameter models at three geographic scales for black spruce in Ontario, Northern Journal of Applied Forestry, 21(2): 83–92.
- Penner, M., Power, C., Muhairve, C., Tellier, R., & Wang, Y. 1997. Canada's forest biomass resources: deriving estimates from Canada's forest inventory. Information Report BC-X-370, Canadian Forest Service, 33 p, Pacific Forestry Centre, Victoria, Canada.
- Shen, T., Chao, A., & Lin, F. 2003. Predicting the number of new species in further taxonomic sampling, Ecology, 84(3): 798–804.

- Temesgen, H., & Gadow, K. 2004. Generalized heightdimater models—an application for major tree species in complex stands of interior British Columbia, European Journal of Forest Research, 123(1): 45–51.
- Temesgen, H., Hann, D., & Monleon, V. 2007. Regional height-diameter equations for major tree species of southwest Oregon, Western Journal of Applied Forestry, 22(3): 213–219, 2007.
- Toivonen, H., Mannila, H., Korhola, A., & Olander, H. 2001. Applying Bayesian statistics to organismbased environmental reconstruction, Ecological Applications, 11(2): 618–630.
- Van Laar, A., & Akca, A. 2007. Forest Mensuration. Springer, Dordrecht, 384 p, The Netherlands.
- Wyckoff, P., & Clark, J. 2000. Predicting tree mortality from diameter growth: a comparison of maximum likelihood and Bayesian approaches, Canadian Journal of Forest Research, 30(1): 156–167.
- Zapata, M., Sierra, C., & Alleman, L. 2012. Probability distribution of allometric coefficients and Bayesian estimation of aboveground tree biomass, Forest Ecology and Management, 277: 173–179.
- Zhang, Z., Duan, A., Zhang, J., & Xiang, C. 2014. Estimating tree height—diameter models with the Bayesian Method, The Scientific World Journal: 1–9.